

# General copula-checking strategies for bivariate failures

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WNAR

# Background

- Interest: Familial aggregation of disease
- Outcome: age at onset  
*unaffected subjects censored*  
*acknowledges importance of early onset*
- Aggregation:  
*existence, strength, implications?*

## Bivariate Failure Time

- Data sampled in clusters (families)  
 $i = 1, \dots, n$
- Clusters contain paired failures  $(T_1, T_2)$   
*can be extended beyond pairs*
- Example: Ages of appendectomy  
*Duffy, Australian Twin Study (1990)*
- How to describe, model dependence?

## Bivariate Failure Time

- $\xi_i$ : latent 'frailty' for  $i$ th pair
- $(T_{i1}, T_{i2})$  independent given  $\xi_i$
- Hazard for  $T_{ij}$  given  $\xi_i$

$$\xi_i \lambda_{0j}(t)$$

- $\xi$  belong to parametric family  $g(\cdot; \beta)$
- Choice of  $g(\cdot)$  affects 'dependence'  
*gamma, positive stable, inverse Gaussian*

## Association in Frailty Models

- Frailty unobserved
- Choice of  $g(\cdot)$  reflected by cross-ratio (CR) function.

$$\theta(t_1, t_2) := \frac{\lambda_1(t_1|T_2 = t_2)}{\lambda_1(t_1|T_2 > t_2)} := \frac{\lambda_2(t_2|T_1 = t_1)}{\lambda_2(t_2|T_1 > t_1)}$$

- Relative hazard at  $t_1$   
*given failure history of  $T_2$  at  $t_2$*
- $\theta(\cdot, \cdot) \iff g(\cdot)$   
*Oakes (1989)*

## Selecting a Frailty Model

- Gamma: Constant CR function
- Positive Stable: CR markedly decreasing in time  
*converges to 1*
- Inverse Gaussian: CR decreasing in time  
*converges to some limit*
- Choice of model can lead to different inference

# Checking Methods

A useful method would....

- Be able to check any model
- Unify graphical/inferential methods
- Closely tied to cross-ratio functions

I'll describe such an approach

# A Local Association Measure

	$\text{pr}(T_1 = t_1)$	$\text{pr}(T_1 > t_1)$	
$\text{pr}(T_2 = t_2)$	$a$	$b$	$a + b$
$\text{pr}(T_2 > t_2)$	$c$	$d$	
	$a + c$		$n$

- $\theta(t_1, t_2) := an / (a + b)(a + c)$
- $H_{11}(dt_1, dt_2) = a/n$
- $H_{10}(dt_1, t_2) = (a + c)/n$
- $H_{01}(t_1, dt_2) = (a + b)/n$
- $\theta(t_1, t_2) := H_{11}(dt_1, dt_2) / H_{10}(dt_1, t_2)H_{01}(t_1, dt_2)$

## Overview

- Has Mantel-Haenszel feel
- Cumulative sum of residuals at  $(t_1, t_2)$
- Residuals based on 2 by 2 table
- Residuals  $\propto \hat{\theta}_{np}(t_1, t_2) - \hat{\theta}_p(t_1, t_2)$

*non-model-based - model based*

*observed - expected*

## The Residuals

$\{t_{1(1)}, \dots, t_{1(d_1)}\}$ :  $d_1$  failures of  $T_1$

$\{t_{2(1)}, \dots, t_{2(d_2)}\}$ :  $d_2$  failures of  $T_2$

At failure times, form the 2-by-2 table

	$T_1 = t_{1(i)}$	$T_1 > t_{1(i)}$	
$T_2 = t_{2(j)}$	$\Delta_{ij}$		$n_{2(j)}$
$T_2 > t_{2(j)}$			
	$n_{1(i)}$		$n_{ij}$

# A Checking Process

- $\hat{\theta}_p(\cdot, \cdot)$ : estimated CR function
- The model-checking process:

$$F(t_1, t_2) = \sqrt{n} \int_0^{t_2} \int_0^{t_1} W(s_1, s_2) \{ \hat{H}_{11}(ds_1, ds_2) - \hat{\theta}_p(s_1, s_2) \hat{H}_{10}(ds_1, s_2) \hat{H}_{01}(s_1, ds_2) \}$$

- Process: cumulative sum of residuals
- Residual at  $(s_1, s_2)$ :  $d\hat{H}_{11} - \hat{\theta}_p(s_1, s_2)d\hat{H}_{10}d\hat{H}_{01}$   
*Proportional to the 2 by 2 table residuals*
- $W_n(\cdot, \cdot)$ : arbitrary weight function  
*e.g., size of risk set, estimated surv fn.*

## Three hazard functions

- $H_{11}(\cdot, \cdot), H_{10}(\cdot, \cdot), H_{01}(\cdot, \cdot)$ : bivariate hazards
- Hazards of double and single failures
- Together they specify joint survival function
- Estimates have Aalen-like structure
- Theory for hazards estimates developed  
*Gill, van der Laan, Wellner (1995)*  
*consistency, weak convergence, bootstrap*

## Appealing Features

- Interpretable process
- $F(t_1, t_2)$  increases, CR locally underestimated
- $F(t_1, t_2)$  decreases, CR locally overestimated
- Any  $\hat{\theta}_p(\cdot, \cdot)$  can be used
- Can use some existing theory
- Graphical and testing component

## Under Correct Frailty

- & Regularity on  $\sqrt{n}\{\hat{\theta}_p(t_1, t_2) - \theta_0(t_1, t_2)\}$
- $F(t_1, t_2)$ : approximately mean 0
- $F(t_1, t_2) \implies \mathcal{F}(t_1, t_2)$

$\mathcal{F}(t_1, t_2)$  mean 0, Gaussian process

- $\{F^\sharp(t_1, t_2) - F(t_1, t_2)\} \implies \mathcal{F}(t_1, t_2)$

$F^\sharp$ : bootstrapped process

bootstrap can approximate  $\mathcal{F}(t_1, t_2)$

## Test for Fit

- Consider two simple summaries of  $F(\cdot, \cdot)$
- $Z_\tau = F(\tau_1, \tau_2) / \hat{SE} \{F(\tau_1, \tau_2)\}$   
*compared to  $N(0, 1)$*   
 *$\tau_1, \tau_2$  some late time*  
*bootstrap used to estimate SE*
- $Q = \sup_t |F(t_1, t_2)|$   
*compare to  $Q^\# = \sup |\{F^\# - F\}|$*

## Simulations: Null

- Generate gamma frailty data
- Fit gamma frailty model
- 50% type I censoring
- 50, 100, 200 pairs
- Kendall's  $\tau$  : 0.25,0.50,0.75
- 2,000 datasets, 200 bootstrap samples
- Empirical size of  $F_\tau$  and  $Q$  for gamma frailty

## Size of Tests: Gamma Frailty

$\tau$		$n = 50$	$n = 100$	$n = 200$
0.25	$F_\tau$	0.06	0.07	0.07
	$Q$	0.05	0.06	0.07
0.50	$F_\tau$	0.06	0.06	0.05
	$Q$	0.04	0.04	0.05
0.75	$F_\tau$	0.03	0.05	0.04
	$Q$	0.02	0.03	0.04

## Simulations: Alternative

- Generate positive stable frailty data
- Fit gamma frailty model
- 50% type I censoring
- $n = 50, 100, 200$  pairs
- Kendall's  $\tau : 0.25, 0.50, 0.75$
- 2,000 datasets, 200 bootstrap samples
- Estimated power based on  $F_\tau$  and  $Q$  statistics

## Power to Reject Gamma Frailty

$\tau$		$n = 50$	$n = 100$	$n = 200$
0.25	$F_\tau$	0.34	0.57	0.84
	$Q$	0.35	0.60	0.89
0.50	$F_\tau$	0.74	0.97	1.00
	$Q$	0.73	0.96	1.00
0.75	$F_\tau$	0.81	1.00	1.00
	$Q$	0.70	1.00	1.00

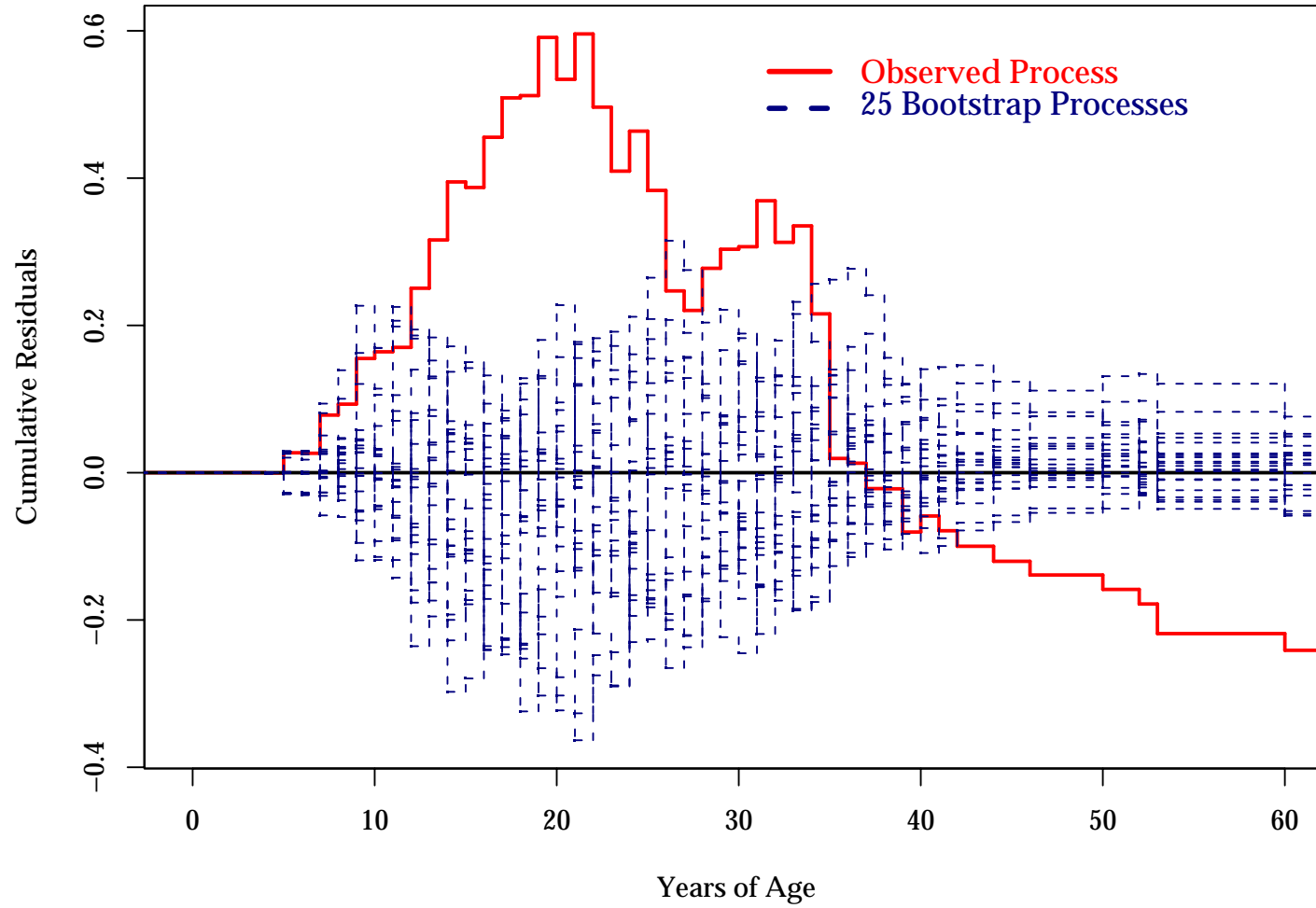
## Australian Twin Study

- 1218 Australian female monozygotic twins
- Age at appendectomy recorded  
*censored at current age*
- Duffy et al (1990) *Am J Hum Genet*
- Strong clustering  
*higher in MZ than DZ twins*
- Gamma Frailty fit:  $\hat{\theta}_p(t_1, t_2) = 2.83$   
*significant clustering*

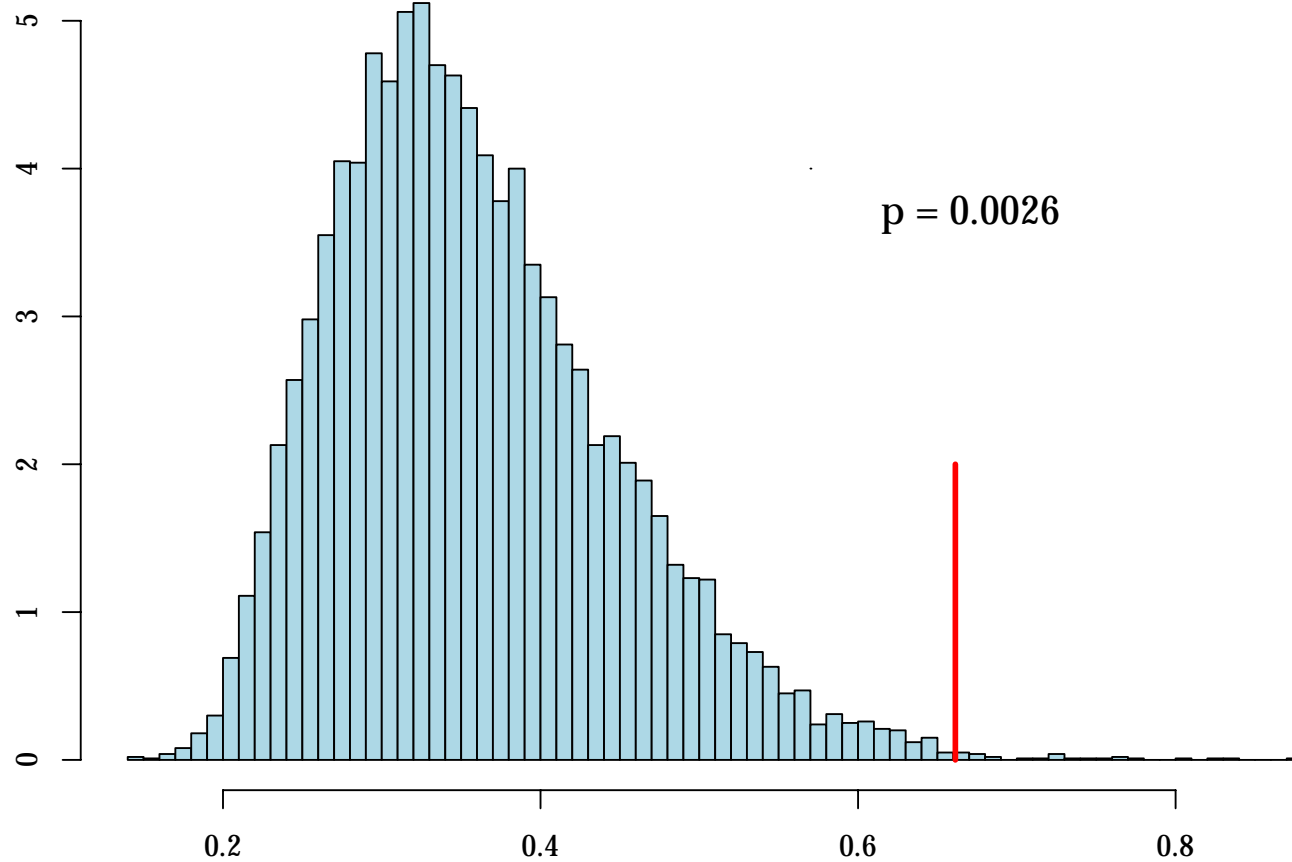
## Constant CR function?

- Non-constant CR function suggested  
*Fan et al, 2000; Kooperberg*
- Glidden (1999): martingale residual method  
*non-significant,  $p = 0.12$*
- Applied method  
*unequivocal rejection of gamma frailty*
- 3-D plot quite overwhelming  
*can graph  $F(t, t), t \in [0, \min(\tau_1, \tau_2)]$*

# Observed and Null Processes

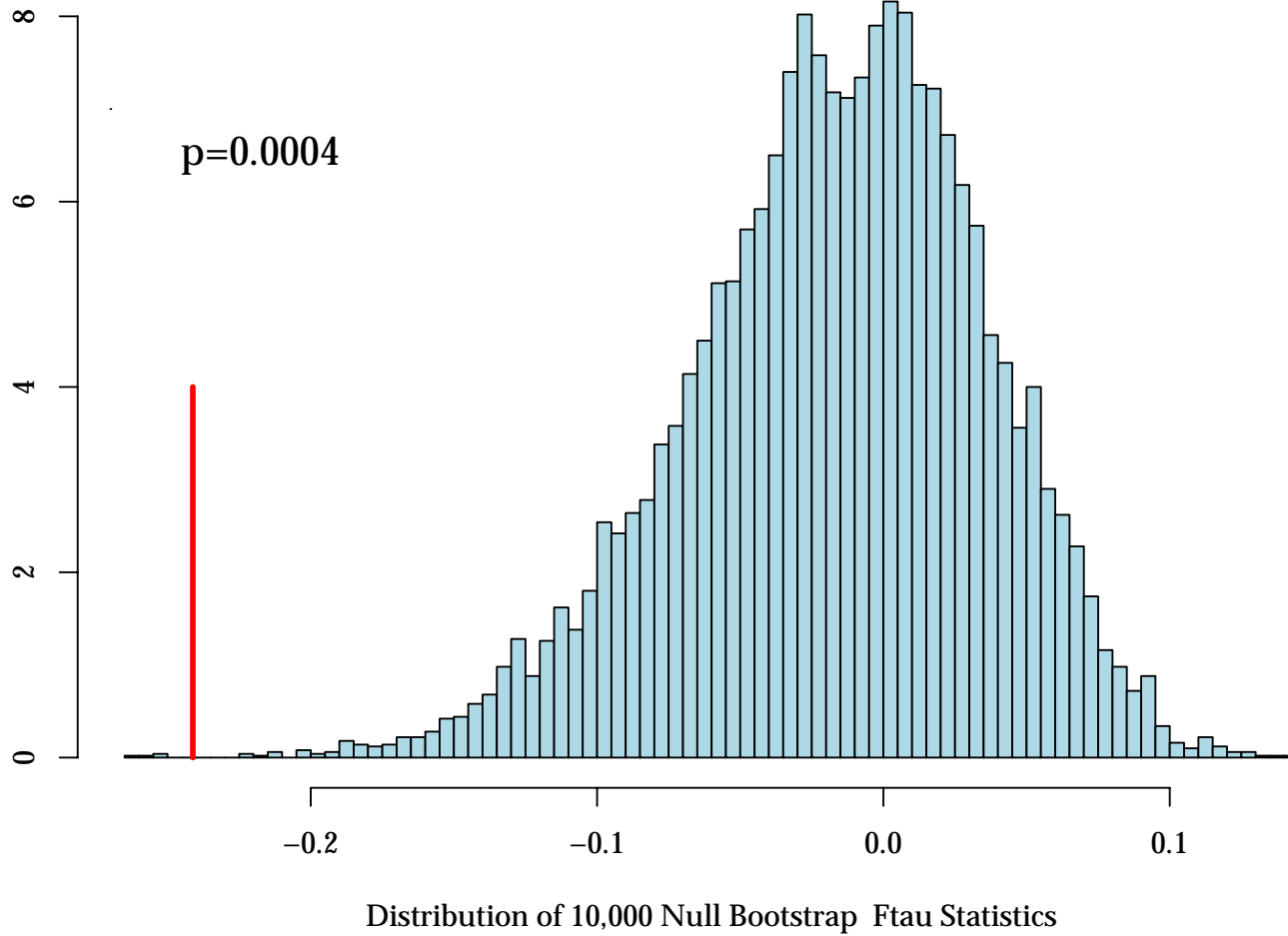


# Observed and Null $Q$ Statistic



Distribution of 10,000 Null Bootstrap Supremum Statistics

# Observed and Null $F_\tau$ Statistic



## Results

- Properties of  $F$  dependent on  $\hat{\theta}_p$   
*no completely general theory*
- Bootstrap work well  
*well-attuned to approximate null distn*
- Good properties, good power
- 3-D plots unwieldy  
*2-D more informative, manageable*
- $F_\tau$  powerful for monotone alternatives  
*e.g. gamma vs. positive stable*